

Take-away from "Intuition about Fisher Information": ①

$$\frac{d^2}{d\theta} \ell(\theta | x_1, \dots, x_n) \Big|_{\theta = \hat{\theta}^{MLE}} \text{ large in magnitude}$$

$\Leftrightarrow \ell(\theta | x_1, \dots, x_n)$ very curved at $\hat{\theta}^{MLE}$

$\Leftrightarrow \ell(\theta | x_1, \dots, x_n)$ changes quickly as we move away from $\hat{\theta}^{MLE}$

\Leftrightarrow ~~data~~ data provide a lot of "information" that $\hat{\theta}^{MLE}$ is better than other possibilities

Consider the form of $\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n)$:

$$\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n) = \frac{d^2}{d\theta^2} \log [L(\theta | x_1, \dots, x_n)]$$

$$= \frac{d^2}{d\theta^2} \log \left[\prod_{i=1}^n f(x_i | \theta) \right]$$

$$= \frac{d^2}{d\theta^2} \sum_{i=1}^n \log \{ f(x_i | \theta) \}$$

$$= \sum_{i=1}^n \frac{d^2}{d\theta^2} \log \{ f(x_i | \theta) \}$$

overall magnitude of $\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n)$

grows with the sample size.

- Subset 1 had a sample size of 56, more information about λ .
- Subset 2 had $n = 4$, less information about λ .

Note: $\frac{d^2}{d\theta^2} \ell(\hat{\theta}^{MLE} | x_1, \dots, x_n) < 0$ ($\hat{\theta}^{MLE}$ maximizes $\ell(\theta)$) ②

$-\frac{d^2}{d\theta^2} \ell(\hat{\theta}^{MLE} | x_1, \dots, x_n) > 0$ has a more intuitive sign:
larger value \Leftrightarrow more information.

Def. (to be refined): The observed Fisher information about a parameter θ is

$$J(\theta^*) = -\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n) \Big|_{\theta = \theta^*}$$

Note: Most often evaluated at the MLE:

$$J(\hat{\theta}^{MLE}) = -\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n) \Big|_{\theta = \hat{\theta}^{MLE}}$$

The observed Fisher information from one observation $X_i = x_i$ is

$$J_i(\theta^*) = -\frac{d^2}{d\theta^2} \ell(\theta | x_i) \Big|_{\theta = \theta^*} = \frac{d^2}{d\theta^2} \log \{ f_{X_i}(x_i | \theta) \} \Big|_{\theta = \theta^*}$$

If observations independent, $J(\theta^*) = \sum_{i=1}^n J_i(\theta^*)$

Ex.: Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Poisson}(\lambda)$.

$$E(X_i) = \lambda, \text{ MLE is } \hat{\lambda}^{MLE} = \bar{X}, f_{X_i}(x_i | \lambda) = e^{-\lambda} \cdot \frac{\lambda^{x_i}}{x_i!}$$

Find the observed Fisher information about λ

if we observe x_1, \dots, x_n ,

$$J_i(\lambda^*) = -\frac{d^2}{d\lambda^2} \log \{ f_{X_i}(x_i | \lambda) \} \Big|_{\lambda = \lambda^*} = -\frac{d^2}{d\lambda^2} \log \left[e^{-\lambda} \cdot \frac{\lambda^{x_i}}{x_i!} \right] \Big|_{\lambda = \lambda^*}$$

$$= -\frac{d^2}{d\lambda^2} \left[-\lambda + x_i \log(\lambda) - \log(x_i!) \right] \Big|_{\lambda = \lambda^*}$$

$$= -\frac{d}{d\lambda} \left[-1 + \frac{x_i}{\lambda} \right] \Big|_{\lambda = \lambda^*}$$

$$= \frac{x_i}{\lambda^2}$$

$$J(\lambda) = \sum_{i=1}^n J_i(\lambda) = \sum_{i=1}^n \frac{x_i}{\lambda^2} = \frac{1}{\lambda^2} \cdot \sum_{i=1}^n x_i = \frac{1}{\lambda^2} \cdot n \cdot \frac{1}{n} \sum_{i=1}^n x_i = \frac{n}{\lambda^2} \bar{x}$$

Evaluate at $\hat{\lambda}^{MLE} = \bar{x}$:

$$J(\hat{\lambda}^{MLE}) = J(\bar{x}) = \frac{n}{(\bar{x})^2} \cdot \bar{x} = \frac{n}{\bar{x}}$$

Def.: The Fisher information is the expected value of the observed Fisher information.

$$I(\theta^*) = -E \left[\frac{d^2}{d\theta^2} \ell(\theta | x_1, \dots, x_n) \Big|_{\theta=\theta^*} \right]$$

↑ capital X's

• on average across all samples of size n , what is the curvature of the log-likelihood function at the parameter value θ^* ?

Def.: The Fisher information from one observation is

$$I_i(\theta^*) = -E \left[\frac{d^2}{d\theta^2} \ell(\theta | x_i) \Big|_{\theta=\theta^*} \right]$$

↑ this is the closest our book comes to defining Fisher information, p. 276

Note: $I(\theta^*) = \sum_{i=1}^n I_i(\theta^*)$

Actual Definition of Fisher information

$$I(\theta^*) = E \left[\left\{ \frac{d}{d\theta} \ell(\theta | x_1, \dots, x_n) \right\}^2 \right] = E \left[\left\{ \frac{d}{d\theta} \sum_{i=1}^n \log(f_X(x_i | \theta)) \right\}^2 \right]$$

Thm.: If $f_X(x | \theta)$ satisfies "regularity conditions"
 ↳ twice continuously differentiable, support of f_X doesn't depend on θ , other stuff, ...
 then the two definitions are equivalent:

$$E \left[\left\{ \frac{d}{d\theta} \ell(\theta | x_1, \dots, x_n) \right\}^2 \right] = -E \left[\frac{d^2}{d\theta^2} \ell(\theta | x_i) \Big|_{\theta=\theta^*} \right]$$

Pf: Enough to show for $n=1$.

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Two preliminary claims:

$$1) \frac{d}{d\theta} \int_{-\infty}^{\infty} f_x(x|\theta) dx = \frac{d}{d\theta} 1 = 0$$

$$2) \frac{d}{d\theta} \log \{f_x(x|\theta)\} = \frac{1}{f_x(x|\theta)} \cdot \frac{d}{d\theta} f_x(x|\theta)$$

$$\Rightarrow \frac{d}{d\theta} f_x(x|\theta) = \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right] \cdot f_x(x|\theta)$$

We get:

$$0 = \frac{d}{d\theta} \int_{-\infty}^{\infty} f_x(x|\theta) dx \quad (\text{claim 1})$$

$$= \int \frac{d}{d\theta} f_x(x|\theta) dx \quad (\text{if regularity conditions satisfied})$$

$$= \int \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right] \cdot f_x(x|\theta) dx \quad (\text{claim 2})$$

Now take second derivative wrt θ :

$$0 = \frac{d}{d\theta} \int \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right] \cdot f_x(x|\theta) dx$$

$$= \int \frac{d}{d\theta} \left(\left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right] \cdot f_x(x|\theta) \right) dx \quad (\text{regularity conditions})$$

$$= \int \left[\frac{d^2}{d\theta^2} \log \{f_x(x|\theta)\} \right] f_x(x|\theta) dx + \int \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right] \cdot \left(\frac{d}{d\theta} f_x(x|\theta) \right) dx$$

(product rule)

$$= \int \left[\frac{d^2}{d\theta^2} \log \{f_x(x|\theta)\} \right] f_x(x|\theta) dx + \int \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right]^2 \cdot f_x(x|\theta) dx$$

(claim 2)

Rearrange:

$$\int \left[\frac{d}{d\theta} \log \{f_x(x|\theta)\} \right]^2 \cdot f_x(x|\theta) dx = - \int \left[\frac{d^2}{d\theta^2} \log \{f_x(x|\theta)\} \right] f_x(x|\theta) dx$$

$$\Leftrightarrow E \left[\left[\frac{d}{d\theta} \ell(\theta|X) \right]^2 \right] = - E \left[\frac{d^2}{d\theta^2} \ell(\theta|X) \right] \quad \therefore$$