

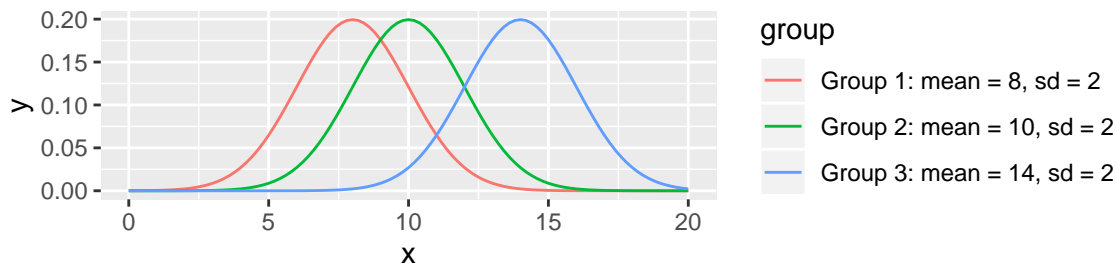
# ANOVA: Concepts for t-based and F-based inference

Sleuth3 Sections 6.2 and 5.2

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## The ANOVA Model

- Observations in group  $i$  follow a Normal( $\mu_i, \sigma^2$ ) distribution
  - (Potentially) different mean for each group
  - Same variance across all groups



## Notation

- We have  $I$  groups ( $I = 3$  for iris example)
- Sample size of  $n_i$  for group  $i$ , total sample size  $n = n_1 + n_2 + \dots + n_I$
- $y_{ij}$ : response variable value for unit  $j$  in group  $i$ 
  - $i$ : which group? ( $i = 1, 2,$  or  $3$  for iris flowers since there are  $I = 3$  species)
  - $j$ : which observational unit within its group? (if  $i = 2$  and  $j = 3$ , we're talking about the 3rd versicolor flower)
- $\bar{y}_i$ : sample mean for group  $i$

## Two Types of Hypotheses:

1.  $H_0 : C_1\mu_1 + C_2\mu_2 + \dots + C_I\mu_I = 0$ .
  - Some combination of means is equal to 0
  - Can specify with one = sign
  - Use a t test
2.  $H_0 : \mu_1 = \mu_2 = \mu_3$ .
  - Some of the group means are actually equal to each other
  - Need multiple = signs to specify
  - Use an F test

t statistic for ANOVA model

	General Set Up	Single Mean	ANOVA (I groups)
<b>Parameter</b>	A number describing the population we are interested in	$\mu$ : population mean (or difference in means with paired data).	$\gamma$ : linear combination of population means for different groups $\gamma = C_1\mu_1 + C_2\mu_2 + \dots + C_I\mu_I$
<b>Estimate</b>	An estimate of the parameter based on the data in our sample	$\hat{\mu} = \bar{Y}$ : sample mean (or difference in sample means with paired data).	Linear combination of sample means for different groups $\hat{\gamma} = C_1\bar{Y}_1 + C_2\bar{Y}_2 + \dots + C_I\bar{Y}_I$
<b>SD(Estimate)</b>	Measures variability of the estimate across different samples.	$\sigma/\sqrt{n}$	$\sigma\sqrt{\frac{C_1^2}{n_1} + \frac{C_2^2}{n_2} + \dots + \frac{C_I^2}{n_I}}$
<b>SE(Estimate)</b>	An estimate of SD(Estimate)	$s/\sqrt{n}$	$s_{pooled}\sqrt{\frac{C_1^2}{n_1} + \frac{C_2^2}{n_2} + \dots + \frac{C_I^2}{n_I}}$
<b>Estimate of <math>\sigma</math></b>	How do we estimate the variance of residuals?	Based on squared differences from the overall sample mean $s = \sqrt{\frac{\sum_{j=1}^n (y_j - \bar{y})^2}{n - 1}}$	Based on squared differences from the group means $s_{pooled} = \sqrt{\frac{\sum_{i=1}^I \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2}{n - I}}$
<b>t statistic</b>	$t = \frac{\text{Estimate} - \text{Parameter}}{\text{SE(Estimate)}}$	$t = \frac{\bar{Y} - \mu}{s/\sqrt{n}}$	$t = \frac{\hat{\gamma} - \gamma}{s_p\sqrt{\frac{C_1^2}{n_1} + \frac{C_2^2}{n_2} + \dots + \frac{C_I^2}{n_I}}}$
<b>Degrees of Freedom</b>		$n - 1$	$n - I$
<b>Confidence Interval</b>	<b>Estimate</b> $\pm t^*SE(\text{Estimate})$	$\bar{Y} \pm t^*SE(\bar{Y})$	$\hat{\gamma} \pm t^*SE(\hat{\gamma})$
<b>P-value</b>	<ul style="list-style-type: none"> <li>• Calculate the t statistic as above, assuming <math>H_0</math> is true (plug in the value of the parameter from <math>H_0</math>)</li> <li>• If the null hypothesis were true, what proportion of samples would have a t statistic at least as extreme as the value you just calculated?</li> </ul>		

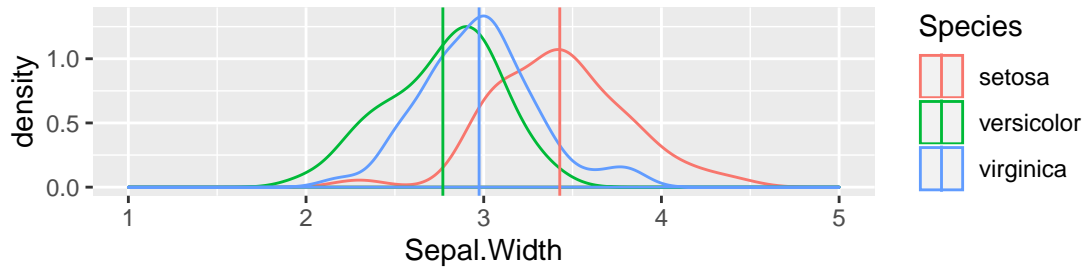
## F tests

Suppose we are conducting a test of  $H_0 : \mu_1 = \mu_2 = \mu_3$  vs.  $H_A$  : at least one of the means differs from the others

We frame this as a comparison of two models.

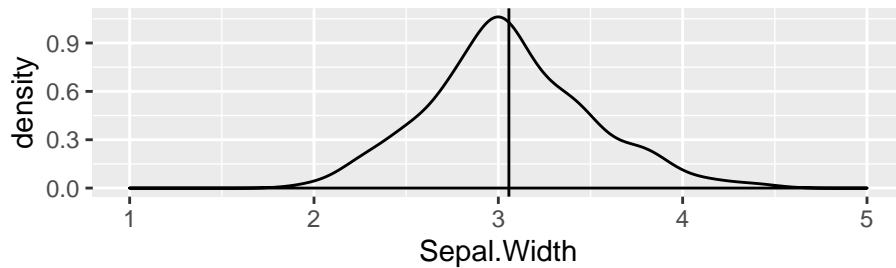
### 1. Full Model, separate means for all groups (corresponds to $H_A$ )

3 mean parameters:  $\mu_1, \mu_2, \mu_3$



### 2. Reduced Model, one mean common to all observations (corresponds to $H_0$ )

1 mean parameter:  $\mu$



### How should we measure the usefulness of a model?

- Suppose we know a flower is from the setosa species, and we want to guess its sepal width. Which guess is better? Why? Can you think of a quantitative way to explain?
  - The group mean for setosa flowers,  $\bar{Y}_1$ . Location of red line in top plot, about 3.5
  - The overall mean for iris flowers,  $\bar{Y}$ . Location of black line in lower plot, about 3.

## Residuals

- **Residual:** difference between observed value for response variable and fitted value for response variable.

$$res_{ij} = Y_{ij} - \bar{Y}_i$$

- In general: Better Model  $\Leftrightarrow$  Better Guesses  $\Leftrightarrow$  Smaller Residuals
- The Full Model will have smaller residuals (on average) than the Reduced Model
- F test answers: are the residuals from the full model enough smaller than the residuals from the reduced model that I think the full model is necessary?

## Measuring the size of residuals from a model

- Residual Sum of Squares: Square the residuals and add them up

$$\sum_i \sum_j (res_{ij})^2 = \sum_i \sum_j (Y_{ij} - \bar{Y}_i)^2$$

- Mean Squared Residual:

$$\frac{\text{Residual Sum of Squares}}{\text{Degrees of Freedom}}$$

## Example

Suppose I have just 3 flowers of each species. Below is an example of the calculation of the RSS for the reduced model (one mean) and the full model (separate means for each group).

i	j	Species	Sepal Width ( $Y_{ij}$ )	Reduced Model			Full Model		
				Mean	Residual	Squared Residual	Mean	Residual	Squared Residual
1	1	setosa	3.9	3.044	0.856	0.733	3.4	0.5	0.25
1	2	setosa	3.1	3.044	0.056	0.003	3.4	-0.3	0.09
1	3	setosa	3.2	3.044	0.156	0.024	3.4	-0.2	0.04
2	1	versicolor	2.4	3.044	-0.644	0.415	2.467	-0.067	0.004
2	2	versicolor	2.6	3.044	-0.444	0.197	2.467	0.133	0.018
2	3	versicolor	2.4	3.044	-0.644	0.415	2.467	-0.067	0.004
3	1	virginica	3.3	3.044	0.256	0.066	3.267	0.033	0.001
3	2	virginica	2.7	3.044	-0.344	0.118	3.267	-0.567	0.321
3	3	virginica	3.8	3.044	0.756	0.572	3.267	0.533	0.284
Total							2.543		1.012

## Extra Sum of Squares

$$\begin{aligned}\text{Extra Sum of Squares} &= \text{Residual Sum of Squares, Reduced Model} - \text{Residual Sum of Squares, Full Model} \\ &= 2.543 - 1.012 \\ &= 1.531\end{aligned}$$

- Always positive because
  - Reduced Model is more limited than Full Model
  - Reduced Model has larger residuals than Full Model
- If Extra Sum of Squares is really big, the Full Model is much better than the Reduced Model
- You can calculate the degrees of freedom for the Extra Sum of Squares in either of two ways:
  - difference in degrees of freedom for the full model and the reduced model:  $(n - 1) - (n - I) = I - 1 = 3 - 1 = 2$
  - difference in number of parameters for the mean between full and reduced model:  $3 - 1 = 2$

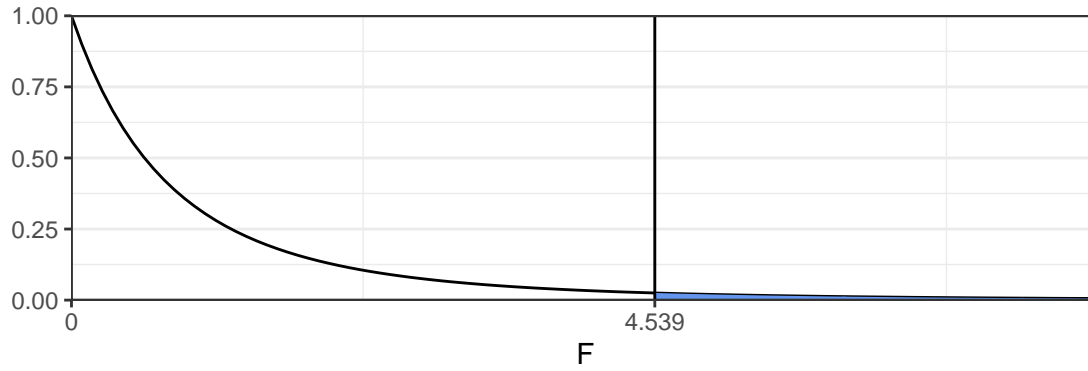
## F Statistic

- “How big is the improvement in Residual Sum of Squares from using the Full Model instead of the Reduced Model”?
  - Size of improvement is measured relative to the size of residuals in the full model

$$\begin{aligned}F &= \frac{(\text{Extra Sum of Squares})/(\text{Extra Degrees of Freedom})}{(\text{Residual Sum of Squares, Full Model})/(\text{Degrees of Freedom, Full Model})} \\ &= \frac{1.531/(3 - 1)}{1.012/(9 - 3)} \\ &= 4.539\end{aligned}$$

- If  $H_0 : \mu_1 = \mu_2 = \mu_3$  is **true**, then...
  - Full Model **isn't better** than Reduced Model
  - Residual Sum of Squares, Full Model is **similar to** Residual Sum of Squares, Reduced Model
  - Extra Sum of Squares is **small**
  - F Statistic is **small**
- If  $H_O : \mu_1 = \mu_2 = \mu_3$  is **not true**, then...
  - Full Model **is better** than Reduced Model
  - Residual Sum of Squares, Full Model is **smaller than** Residual Sum of Squares, Reduced Model
  - Extra Sum of Squares is **large**
  - F Statistic is **large**
- **A large value of F statistic is evidence against  $H_0$**

- For finding p-values we are interested in the probability of getting an F statistic at least as large as the F statistic we got from our sample, if  $H_0$  is true.



- We have to keep track of two degrees of freedom